

# Immigration, Enterprises, and Employment in the European Union

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October 1, 2012

## Abstract

We study the effects of ethnic diversity, measured by the share of immigrants in the total population in the EU member countries, on the number of establishments and employment in the EU. We distinguish between different industries and between different groups of source (migrant-sending) countries; i.e., Eastern European Countries (EECs) and Mediterranean Countries (MPCs). We use a panel data that covers the period 1988-2010, and find that migration from MPCs to the EU has a positive impact on both the number of enterprises and employment, especially in light manufacturing industries. Also migration from MPCs to the EU positively affects employment in construction and heavy manufacturing industries. Similarly, migration from EECs to the EU positively affects employment, especially in food and beverages industries.

**Keywords:** Migration; Enterprises; Employment.

**JEL Classification:** J2; J61; R23.

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# 1 Introduction

Immigration is at the forefront of the European Union's (EU) attention as it is believed it significantly affects economic outcomes for natives through various channels. In this paper, we study the effects of ethnic diversity, measured by the share of immigrants in the total population in the EU member countries, on industry-level employment, and on the variety of products that both natives and immigrants are provided with in these countries. In particular, we are interested in finding out about whether and how the composition of businesses in the EU changes with the influx of immigrants. For example, do we see a change at the *product* extensive margin, such that new products are made available in the market, especially to cater to immigrants' demand for some ethnic/diversified goods, or at the *product* intensive margin such that the share of establishments providing more standardized products increase?

The empirical motivation of our paper is obvious as statistical evidence shows that the immigrant population in the EU is significantly large. At the end of the 1990s, 3.5 per cent of the EU's population (18 million) was of immigrant origin; see Aubarell and Aragall (2005). In about a decade, this number has almost doubled. According to the *News Release* by EUROSTAT (2010) — the statistical office of the EU — at the end of 2008, there were 31.9 million foreign citizens living in the EU, of which 20 million were citizens of countries outside the EU. The share of the EU population that is foreign born is currently estimated at around 10 per cent; see EMPL (2011). Not surprisingly, the most populated five EU Member States (Germany, France, Italy, Spain, and the United Kingdom) — comprising approximately two-thirds of the total EU population — have the highest numbers of foreign-born persons, in absolute terms, the total number corresponding to over 75 per cent of the total immigrant population in the EU; see EUROSTAT (2011a). As is discussed in detail, in the following section, not only may immigrants bring in their knowledge of producing some diversified goods, or make trading such goods possible/less costly, but also they may create significant demand for such goods. So we may eventually see some immigration-triggered changes in consumption and production patterns, especially in countries receiving sufficiently large numbers of foreign-born persons.

In general, people move across countries for several reasons. In particular, employment-related reasons are reported as the main motive behind immigration, although migrants

tend to have low levels of income, and/or are exposed to a higher risk of unemployment <sup>1</sup>, or are likely to be employed in jobs below their educational qualifications. We shall note that there are some important factors contributing to immigrants' such employment experiences, such as the non-recognition of migrants' qualifications and skills which are earned abroad, language barriers, or discrimination, etc.; see EUROSTAT (2011a) and EMPL (2011) for details. These factors may also explain, to some extent, the sectoral distribution of immigrants in the EU Member States.

According to the EU-LFS 2009 data reported by EMPL (2011), immigrants are, generally, under-represented in occupations (i) that require proficiency in the host country language such as office works as they cannot compete with a larger group of native speakers, and (ii) that require high skills/education as in extra-territorial organizations, and education and health sectors, etc. Also they are not well represented in manufacturing, and wholesale and retail trade industries, although there is considerable heterogeneity across countries. On the contrary, they are over-represented in occupations (i) whose demand for skill is sufficiently low such as service sector industries (e.g., hotel and food services, and administrative and support service activities, etc.), and (ii) where the employer is the household (i.e., the household sector that consists in domestic helpers, cleaners and launderers, and personal care workers). Also they are well represented in the construction sector, although as in manufacturing, and wholesale and retail trade industries, the share of immigrant employment in the construction sector shows significant heterogeneity across countries. <sup>2</sup>

In this study, we scrutinize mainly the demand-related impact of the influx of immigrants on the variety of consumption goods available in the host countries. Hence the retail industry, which involves activities that are related to selling goods and services directly to consumers, is given a special emphasis. According to the EU-27, 2008-data, published by Eurostat (2011b), the retail industry is a subgroup of the distributive trades sector, which involves mostly activities that are related to the purchase and resale of goods in the same condition. The distributive trades sector includes 6.1 million enterprises — nearly 30 per cent of the total number of enterprises in the EU non-financial business economy

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<sup>1</sup>Irrespective of the level of education, the unemployment rates of foreign-born persons were systematically higher than for native-born persons, and especially in 2008, this was true in almost all Member States for which data were available (EUROSTAT 2011a: 41).

<sup>2</sup>For a detailed analysis of the sectoral distribution of the share of immigrant employment, see EMPL (2011).

— a large number of which is micro, small, or medium-sized enterprises,<sup>3</sup> and provides employment for almost a quarter of the EU non-financial business economy workforce (32.8 million persons),<sup>4</sup> so it is the largest sector in terms of the number of enterprises, and is almost as large as the manufacturing industry in terms of the number of persons employed.<sup>5</sup> Moreover, the share of the retail industry in total distributive trades is the largest both in terms of the number of enterprises (60 per cent) and of persons employed (55 per cent); see Eurostat statistics, *European Business* (NACE divisions).

We can distinguish between different types of enterprises. In general, establishments that are affiliated with a large firm, which consists of several stores (e.g., chain stores), (i) have complex distribution and inventory control systems, (ii) benefit significantly from scale and scope economies, and (iii) tend to provide more standardized products and offer lower prices. Small, owner-operated/stand-alone stores, however, tend to offer more customized products, and charge higher prices; see Dinlersoz (2004) for details. A positive relationship between the number of small, owner-operated/stand-alone stores and the share of immigrant population, hence, can be associated with the change of the composition of businesses at the product extensive margin, and so with increased diversity of consumption choices. By the same token, if immigrants have higher price elasticities of demand, or if they tend to consume products offered by chain stores, we may well observe a shift of the composition of businesses in the opposite direction.

## 2 Review of the related literature

There is an extensive literature studying potential impacts of immigration in different contexts. One strand of this literature, for example, focuses on the labor-market consequences of immigration, such as whether immigration leads to higher unemployment among natives, especially by crowding out native workers, and whether immigration decreases wages/earnings of native workers, etc. Although the vast majority of research has mainly analyzed the United States (US)<sup>6</sup>, there is a growing and recent literature

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<sup>3</sup>Micro, small, or medium-sized enterprises (SMEs), which comprise 99.8 per cent of all active enterprises in the EU non-financial business economy are mainly concentrated in the distributive trades sector.

<sup>4</sup>The number of persons employed in the non-financial business economy is estimated at about 136.3 million, that is approximately 60 per cent of total employment in the EU.

<sup>5</sup>Manufacturing is the largest sector within the EU non-financial business economy, both in terms of the number of persons employed (33 million) and of value added.

<sup>6</sup>See Hanson (2009) for discussions of this literature.

studying different EU Member States.<sup>7</sup> Much of this literature is indirectly related to our study as we particularly focus on the immigration-induced changes in product diversity. It is, however, worth noting that, as far as the EU Member States are concerned, in most cases, immigrants do not crowd out native workers — since they mostly complement natives in the labor market — nor do they have a significant negative impact on native workers' wages/earnings, which may have indirectly affected consumption choices; see Kerr and Kerr (2011), Münz *et al.* (2007), ILO (2010), UNECE (2002), and references therein, for details. To the contrary, migrant workers contribute to job creation in several ways, ranging from entrepreneurship to increasing domestic demand for goods and services (ILO 2010: 60).

Immigrants generally create social networks in the country that they have settled (OECD 2007). Such networks enable immigrants to opt for self-employment, and so to establish micro, small, or even medium-sized enterprises,<sup>8, 9</sup> which are mostly found in the catering industry, services, and retail trade. Immigrant entrepreneurs that are active in such sectors often provide goods and services that are different from those provided by native entrepreneurs, implying that they may well contribute to the diversity of consumption choices (EC 2006, EMN 2005, ILO 2010). Immigrants may also play a crucial role in facilitating trade through a number of mechanisms as they are linked to both their home and host countries by networks; see Gaston and Nelson (2011), Globerman (1995), and Head and Ries (1998) for details. As argued by Head and Ries (1998), immigrants may have superior knowledge of market opportunities, and so in the presence of transaction costs, they may act as trade intermediaries, and may reduce costs, especially associated with foreign trade.<sup>10</sup> Such costs tend to be significantly high, especially when economic,

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<sup>7</sup>A survey of the main findings of such studies can be found in UNECE (2002), the United Nations Economic Commission for Europe.

<sup>8</sup>According to the European Commission (EC) publication, EC (2006), in Italy, there are some 168,000 such enterprises. In Belgium, in the Brussels area alone, self-employed persons originating from ethnic minority communities are estimated at around 18,000, while for the Flemish region, the number is estimated at about 10,000. In Germany, in 2003, there were 142,000 self-employed non-EU citizens, and in Netherlands, in 2004, 58,000 ethnic entrepreneurs were recorded (p.17).

<sup>9</sup>Among different motives, *immigrant entrepreneurship* is a way to circumvent unemployment, especially given their difficulties in finding paid-employment via formal routes; see e.g., van Delft *et al.* (2000), Constant *et al.* (2005), EMN (2005), and OECD (2007).

<sup>10</sup>This is referred to as the *information bridge hypothesis*, according to which immigrants may have superior knowledge of both the home and host country markets, languages, business practices, laws, and special distribution channels, etc., that may help overcome uncertainty stemming from economic and cultural differences, and differences in political environments across countries. Also immigrants may help reduce economic inefficiencies, which may arise especially due to asymmetric information and incomplete enforcement of contracts; see Dunlevy (2006), and Gaston and Nelson (2011).

cultural, and institutional differences across countries are significant, and when such countries trade specialized and/or differentiated goods. Therefore, immigrants may positively affect trading differentiated goods, which may lead to increased variety of consumption goods in the host country.

There is a sizeable literature on the relationship between immigration and trade. Empirical evidence from this literature, which mainly employs gravity-based estimation techniques, suggests that immigration has indeed a significant positive effect on both exports and imports, and the effect appears to be stronger for imports and for specialized/differentiated goods.<sup>11</sup> This latter finding implies that immigrants may also change the number of varieties of goods available in the host country, especially through their demand/consumption patterns. The idea here is simple. If immigrants have preferences for certain goods produced in their country-of-origin — which may not be available in the country that they immigrate — and if their demand for such goods is sufficiently large — which is likely to occur in countries where the share of immigrants in the total population is sufficiently large — then they may lead the host country to import such differentiated goods.<sup>12</sup> By the same token, immigrants may have a comparative advantage in producing such goods, with which supply may increase. Though a similar effect may stem from any kind of frictions or preferences leading immigrants to increase labor supply in industries producing such differentiated goods.

In this paper, we also study such preference effects of immigrants, but with a different focus. In contrast to the trade literature mentioned above, we want to delineate how the composition of businesses are linked to the share of immigrants in the total population, which is, surprisingly, a far less studied question in the existing literature, and so with which we would like to contribute to the literature. Our paper is closely related to Mazzolari and Neumark (2011) studying the impact of immigration on the diversity of consumption choices. In particular, they try to explain the changes in the number of establishments of different size with the changes in the share of immigrants in the total population. They use establishment-level data for California between 1992 and 2002, and focus on the retail sector and the restaurant sector, the latter of which is given a special emphasis. They find that immigration is associated with fewer stand-alone retail stores,

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<sup>11</sup>See Wagner *et al.* (2002), Peri and Requena-Silvente (2010), and Gaston and Nelson (2011), and references therein, for surveys and discussions of the main findings of this literature.

<sup>12</sup>This preference effect is referred to as the *transplanted home bias* effect as migrants develop tastes before migrating to a country, and as such tastes affect their consumption patterns in the country they immigrate; see White (2007) for discussions of such preference effects.

and a greater number of chains/big-box retailers, which appears to be contradicting with the diversity-enhancing effect of immigration. Although they find a positive relationship between immigration and ethnic diversity in the restaurant sector, which — as they argue — also may stem from comparative advantage of immigrants in the production of ethnic food from their country-of-origin. To the contrary, Olney (2011) argues that the relationship between immigration and the number and size of establishments is mainly driven by firms' relocating their production activities, rather than by immigrants' consumption patterns. He uses a data set that covers 192 U.S. Metropolitan Statistical Areas for the period 1998-2004, and shows that firms respond to immigration both at the extensive margin, which is captured by the net birth rate of establishments, and at the intensive margin, which is captured by the net expansion rate of establishments. According to his results, both the net birth rate and the net expansion rate of establishments increase, especially with low-skilled immigration, the impact of which appears to be much weaker in the non-mobile industries, such as agriculture, mining, and retail trade, than in the mobile industries, such as manufacturing, and finance, professional, management, and administration services. That being said, his data do not allow for calculating immigration by industry, which may have been crucial for an analysis focusing on the production-related effects of immigration in different industries as immigrants are not well represented in those so-called mobile sectors.

Another strand of the literature, to which our paper is indirectly related, looks at how prices change with the influx of immigrants. Lach (2007) employs a store-level price data and shows that the unexpected arrival of a large number of immigrants from the former Soviet Union in Israel, in the 1990s, leads to large and significant reductions in prices. This result may well reflect the demand-side effect of immigration, that is, new consumers (immigrants) have high price elasticity and low search costs, especially vis-à-vis the native population. Given composition effects, we may see the arrival of consumers with different characteristics may offset the demand level changes stemming from the increase in the number of consumers. Bodvarsson *et al.* (2008) analyze the effects of immigration from Cuba to Miami, especially after the Mariel Boatlift of 1980, and find positive demand effects, that is, retail sales per capita increased with the influx of Cuban immigrants. Bodvarsson and Van den Berg's (2009) study, which focuses on Hispanic immigration to Dawson County, Nebraska — a uniquely-segmented economy where immigrants work exclusively in an export sector (the meatpacking industry) but consume locally — suggests that immigration can boost local consumer demand. Similarly, Frattini (2008), focus-

ing on immigration inflows in the UK, between 1995 and 2006, shows that the price of low-value and everyday grocery goods increased in the same period.

Our study differs from the existing studies such that not only it considers the supply-side effects of immigration, but that it treats immigrants as potential consumers of differentiated goods. Also our study is not confined to a particular area in a country, or to a single country. We study the EU Member States, which is also a contribution to the literature that mainly focuses on the US. Moreover, we distinguish between different industries and between different groups of source (migrant-sending) countries; i.e., Eastern European Countries (EECs), and Mediterranean Countries (MPCs). By employing static estimation methods, our study suggests that migration from MPCs to the EU has a positive impact on both the number of enterprises and employment, especially in light manufacturing industries. Migration from MPCs to the EU positively affects employment in construction and heavy manufacturing industries. Similarly, migration from EECs to the EU positively affects employment, especially in food and beverages industries. In the following sections, we introduce our methodology and data, and present our results. The last section provides some concluding remarks.

### 3 Methodology

*The number of establishments* equations are of the reduced form, and are derived from Mazzolari and Neumark (2009). Its theoretical roots can be found in Ottaviano and Peri's (2006, 2008) studies, which incorporate *consumption variety* effects into the study of the economic benefits of immigration. In particular, they employ a general equilibrium model for a small open economy where individuals are differentiated in terms of origin, home-born vs foreign-born, and consume two goods, a homogeneous tradable good and a differentiated, local, non-tradable good. Home-born and foreign-born individuals are assumed to be able to produce different varieties of the non-tradable good. In such a model, the non-tradable good can be thought of as a composite basket of local services whose supply particularly benefits from diversity.

We follow the same approach, and attempt to directly study the relationship between immigration and the composition of products available to consumers. In particular, We scrutinize the effects of immigration on product diversity by looking at the industry-level number of enterprises and industry-level employment. We use the following two equations

to estimate the impacts of immigration on the number of enterprises and on employment:

$$\begin{aligned} enter^j &= \alpha^0 + \alpha^1(EEC/Pop) + \alpha^2(MPC/Pop) + \alpha^3(NAV/Pop) + \alpha^4(REN) + \varepsilon, \\ emp^i &= \beta^0 + \beta^1(EEC/Pop) + \beta^2(MPC/Pop) + \beta^3(NAV/Pop) + \varepsilon, \end{aligned}$$

where the variable  $enter^j$  and  $emp^i$  are the number of enterprises and employment in industry  $j$  and  $i$ , respectively, and  $EEC/Pop$ ,  $MPC/Pop$  and  $NAV/Pop$  are the share of immigrants from EECs and from MPCs, and the share of native population in total population, respectively, and  $REN$  is the total remuneration paid to employees. These equations are also estimated by using the changing rates of the variables.

We use a panel-based approach so as to deal with unobserved country-pair heterogeneity, because conventional cross-section estimation techniques fail to model such heterogeneity, and so may yield biased estimates; see Cheng and Wall (2005) and Carrre (2006) for details. Cross-section specifications also fail to properly account for possible omitted variables bias; see De Benedictis and Taglioni (2011) for discussions. The two commonly used panel estimation techniques are the fixed-effects (FE) and the random-effects (RE) estimation methods. The main difference between the two methods is that the FE method allows country-pair individual effects to be correlated with regressors, whereas the RE method assumes that individual effects are uncorrelated with all regressors. As the FE method transforms data into deviations from individual means, ignoring the between-groups variance, it cannot provide estimates for the coefficients of time-invariant regressors such as distance. Although this is a disadvantage, an FE estimator is unbiased and consistent in the presence of correlation between individual effects and regressors, whereas the RE estimator is not. The common practice to choose which model to use is to employ a Hausman specification test, as suggested by Hausman (1978). We follow this strategy such that we first employ both the FE and RE models when estimating the effects of immigration on product diversity, then we employ a Hausman test.

## 4 Data and Results

Our migration data covers the period 1988-2010, and provides information on the number of immigrants in the EU. We distinguish between immigrants in terms their country of birth. So we have immigrants whose home country either belongs to the group of MPCs or to that of EECs. Also we distinguish between immigrants in terms of gen-

der and age. That said, the migration variable used in the econometric estimation includes the total number of immigrants. We extract our data mainly from two data sources, <http://ec.europa.eu/eurostat> from which we collect our migration data, and the OECDs Structural and Demographic Business Statistics from which we collect the enterprises and employment data. We shall note that the latter data source provides information at a very detailed sectoral level, especially on turnover, value-added, production, operating surplus, employment, labor costs and investment. The breakdown by industrial sector, including services, is supplemented by a further breakdown into size classes. The database also includes business demography statistics, such as enterprise birth, death and survival rates, as well as the number of high-growth enterprises and gazelles, especially from 1995 onwards.

We look at the relationship between migration and product diversity, both in terms of employment and the number of enterprises. We consider seven industries: (1) mining and quarrying, (2) food products, beverages and tobacco, (3) light manufacturing, (4) heavy manufacturing, (5) electricity, gas and water supply, (6) construction, (7) wholesale and retail trade, hotels and restaurants. Dependent variables (employment and the number of enterprises) are specified first as levels, then as the rates of change, and independent variables are adjusted accordingly. The FE and the RE models are estimated, and then a Hausman specification test is performed. Therefore, eight models for each industry are estimated and, in total, 56 econometric estimations are carried out. We shall note that, in almost all cases, the RE models are rejected. So we mainly focus on the results of the FE models, although, in some cases, we present also the results of the RE models.

Table 1 summarizes our findings from the estimation of the industry-level number of enterprises. In Table 1, the signs, (+) and (−), mean *positive significant* and *negative significant*, respectively. As for the variables,  $Y1$  stands for the number of enterprises, and  $X1$ ,  $X2$ , and  $X3$  stand for the share of immigrants from EECs and from MPCs, and the share of native population in total population, respectively.  $X4$  is the total remuneration paid to employees. The model is also re-estimated by substituting  $X5$ , the total population including migrants, for  $X3$ , and the results are included in the lower section of the table. As is given by Table 1, immigrants from MPCs have a positive impact on the number of enterprises, especially in light manufacturing industries. That said, immigrants from EECs have a negative impact on the number of enterprises, especially in electricity, gas and water supply industries. As for the impact of total remuneration paid to employees on the number of enterprises, our results from the model in which  $X3$  is considered suggest

	Fixed effect					Random effect				
	Variables					Variables				
	Dep. Var.	X1	X2	X3	X4	Dep. Var.	X1	X2	X3	X4
<b>Industries</b>										
1 Mining and Quarrying	Y1					Y1				-
2 Food Products, Beverages and Tobacco	Y1					Y1				
3 Light Manufacturing	Y1	+				Y1		+		
4 Heavy Manufacturing	Y1					Y1				+
5 Electricity, Gas and Water Supply	Y1			+		Y1				+
6 Construction	Y1			+		Y1				+
7 W/sale, Retail Trade, Hotels and Restaurants	Y1			+		Y1				+

  

	Fixed effect					Random effect					
	Variables					Variables					
	Dep. Var.	X1	X2	X3	X4	Dep. Var.	X1	X2	X3	X4	X5
<b>Industries</b>											
1 Mining and Quarrying	Y1					Y1					+
2 Food Products, Beverages and Tobacco	Y1					Y1					
3 Light Manufacturing	Y1	+				Y1		+			+
4 Heavy Manufacturing	Y1					Y1					+
5 Electricity, Gas and Water Supply	Y1				+	Y1					
6 Construction	Y1				+	Y1					
7 W/sale, Retail Trade, Hotels and Restaurants	Y1				+	Y1					

Table 1: Results – the number of enterprises



	Fixed effect				Random effect			
	Variables				Variables			
	Dep. Var.	X6	X7	X8	Dep. Var.	X6	X7	X8
<b>Industries</b>								
1 Mining and Quarrying	Y2				Y2			
2 Food Products, Beverages and Tobacco	Y2		-		Y2			
3 Light Manufacturing	Y2		-		Y2			
4 Heavy Manufacturing	Y2				Y2			
5 Electricity, Gas and Water Supply	Y2				Y2		+	+
6 Construction	Y2		+		Y2		+	
7 W/sale, Retail Trade, Hotels and Restaurants	Y2				Y2			
	Fixed effect				Random effect			
	Variables				Variables			
	Dep. Var.	X6	X7	X9	Dep. Var.	X6	X7	X9
<b>Industries</b>								
1 Mining and Quarrying	Y2				Y2			
2 Food Products, Beverages and Tobacco	Y2			-	Y2			
3 Light Manufacturing	Y2			-	Y2			
4 Heavy Manufacturing	Y2				Y2			
5 Electricity, Gas and Water Supply	Y2				Y2			+
6 Construction	Y2				Y2			
7 W/sale, Retail Trade, Hotels and Restaurants	Y2				Y2			

Table 3: Results – the rates of change in employment

a positive relationship, especially in construction, wholesale and retail trade, hotels and restaurants, and electricity, gas and water supply industries. Also, in the model in which the total population including migrants is substituted for the share of native people in total population we find that the impact of the total population including migrants on the number of enterprises is negative in food products, beverages and tobacco industries, and is positive in electricity, gas and water supply industries.

We summarize our estimation results for the impact of immigration on industry-level employment in Table 2. As before, in Table 2, the signs, (+) and (−), refer to as *positive significant* and *negative significant*, respectively. As for the variables,  $Y3$  stands for employment, and  $X10$ ,  $X11$ , and  $X12$  stand for the share of immigrants from EECs and from MPCs, and the share of native population in total population, respectively. We also re-estimate the model by substituting  $X13$ , the total population including migrants, for  $X12$ , and present the results in the lower section of the table. As is given by Table 2, immigrants from MPCs have a positive impact on employment in both light and heavy manufacturing industries. That said, immigrants from EECs have a positive impact on employment in food products, beverages and tobacco industries. Similarly, an increase in total population (including migrants) increases employment in wholesale and retail trade, hotels and restaurants, and decreases employment in light and heavy manufacturing industries, and in food products, beverages and tobacco industries.

Finally, we scrutinize how the rates of change in immigration affect industry-level employment, results of which are given by Table 3, where the signs, (+) and (−), refer to as *positive significant* and *negative significant*, respectively. The variable  $Y2$ , now, stands for the rate of change in employment, that is,  $\text{employment}(t - (t - 1))/\text{employment}(t - 1)$ . Similarly, the variables  $X6$ ,  $X7$ , and  $X8$  stand for the change in the share of immigrants from EECs and from MPCs, and the change in the share of native population in total population, respectively. We re-estimate the model by substituting  $X9$ , the change in the total population including migrants, for  $X8$ , and present the results in the lower section of the table. As is consistent with our previous results, in this case, our results suggest a negative relationship between the change in total population and the change in employment, especially in food products, beverages, and tobacco, and light manufacturing industries. Also we observe the same effects in the same industries even when we do not substitute the change in the total population for the change in the share of native population in total population. In this case, we observe a positive impact of immigrants from MPCs on employment, especially in the construction industry.

## 5 Concluding Remarks

In this study, we have scrutinized the effects of ethnic diversity, measured by the share of immigrants in total population in the EU member countries, on the number of establishments and employment in the EU. We have distinguished between different industries and between different groups of source (migrant-sending) countries; i.e., Eastern European Countries (EECs) and Mediterranean Countries (MPCs).

One solid conclusion that can be drawn from our estimation results is that immigration from MPCs certainly has a capacity building effect, especially on manufacturing industries. Empirical evidence is more solid in light manufacturing as immigrants from MPCs have a positive impact both on the number of enterprises, and on employment, whereas for heavy manufacturing, the only positive effect is on employment. In contrast, immigration from EECs has a negative impact on the number of enterprises in electricity, gas and water supply industries, but a positive impact on employment, especially in food products, beverages and tobacco industries. We suppose this distinction between the results regarding employment effects is due to the historical colonial relationship between MPCs and the EU countries. Taking this relationship for granted, and given the large share of immigrants from MPCs in the EU for long time, immigrants from MPCs probably satisfy their needs of ethnic food and beverages more easily, and so do not feel the necessity to expand production capacity of ethnic food and beverages. Another possibility is that there are already sufficient imports of ethnic food and beverages, especially from MPCs to the EU. Therefore less qualified immigrants from MPCs prefer to work in less demanding but better paying positions as in the light and heavy manufacturing industries. As for the less qualified immigrants from EECs, food and beverages might be the only industry that they may get employed, especially in the short run.

Another interesting result is that there is a negative relationship between the change in the share of native population in total population and the change in employment, especially in food, beverages, and tobacco, and light manufacturing industries. Actually, this finding is consistent with the ones that we have already discussed, such that natives probably prefer to have a better work environment and to take up better paying jobs, and so they rather leave low-profile jobs to immigrants, which is also supported by the employment argument above.

Finally, our results have shown that an increase in the change in the share of immigrants

from MPCs increases employment in the construction industry, which might be expected, especially due to low skill levels of immigrants, or due to some other factors that we have already discussed in the introduction section (e.g. the non-recognition of immigrants' qualifications and skills which are earned abroad).

We believe further studies, especially looking at industry-wise trade and migration, are warranted in order to draw a complete picture. It would be interesting to find complementary results, especially between the change in trade patterns and the change in industrial enterprises. Also we leave computing the factor content of industries in order to see the similarities and differences, and factors affecting international migration as future research.

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**Not for publication**

## APPENDIX: RESULTS

### Changes in the number of enterprises

- $Y1 = \# \text{ of enterprises } (t - (t-1)) / \text{enterprises } (t - 1)$   
 $X1 = \text{immigration from EECs } (t - (t-1)) / \text{total population } (t - 1)$   
 $X2 = \text{immigration from MPCs } (t - (t-1)) / \text{total population } (t - 1)$   
 $X3 = \text{native people } (t - (t-1)) / \text{total population } (t - 1)$   
 $X4 = \text{per labor renumeration}$

**Table A1: Mining and quarrying (I)**

```

. xtreg y1 x1 x2 x3 x4, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0269                Obs per group:  min =    9
      between = 0.1711                  avg   =   9.0
      overall  = 0.0001                  max   =    9

corr(u_i, Xb) = -0.6561                F(4, 12)       =   2.01
                                           Prob > F        =   0.1564

                                           (Std. Err. adjusted for 13 clusters in n)

```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-5.826768	4.341554	-1.34	0.204	-15.2862	3.632664
x2	1.659532	25.55738	0.06	0.949	-54.02521	57.34428
x3	-10.20824	6.065679	-1.68	0.118	-23.42422	3.007742
x4	-168.6257	697.3205	-0.24	0.813	-1687.957	1350.705
_cons	.0579442	.0276468	2.10	0.058	-.0022929	.1181814
sigma_u	.05449419					
sigma_e	.13743415					
rho	.135861					(fraction of variance due to u_i)

```

Modified Wald test for groupwise heteroskedasticity
in fixed effect regression model

H0: sigma(i)^2 = sigma^2 for all i

chi2 (13) = 1.0e+05
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data
H0: no first-order autocorrelation
F( 1, 12) = 6.916
Prob > F = 0.0220

```

**Table A2: Mining and quarrying (II) - population to replace X3**

```

. xtreg y1 x1 x2 dlnpop x4, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0268                Obs per group:  min =    9
      between = 0.1709                  avg   =   9.0
      overall  = 0.0001                  max   =    9

corr(u_i, Xb) = -0.6552                F(4, 12)       =   2.00
                                           Prob > F        =   0.1588

                                           (Std. Err. adjusted for 13 clusters in n)

```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	4.373359	6.279704	0.70	0.499	-9.30894	18.05566
x2	11.90253	21.68721	0.55	0.593	-35.34984	59.1549
dlnpop	-10.25921	6.122394	-1.68	0.120	-23.59877	3.080335
x4	-168.3632	697.7702	-0.24	0.813	-1688.674	1351.947
_cons	.0580071	.0278091	2.09	0.059	-.0025838	.118598
sigma_u	.05441326					
sigma_e	.13744015					
rho	.13550214					(fraction of variance due to u_i)

```

Wooldridge test for autocorrelation in panel data
H0: no first-order autocorrelation
F( 1, 12) = 6.913
Prob > F = 0.0220

```

**Table A3: Food products, beverages and tobacco (I)**

```
. xtreg y1 x1 x2 x3 x4, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =    117
Group variable: n                     Number of groups =    13
R-sq:  within = 0.1123                Obs per group:  min =     9
      between = 0.0867                  avg           =    9.0
      overall  = 0.0088                  max           =     9
corr(u_i, Xb) = -0.6946                F(4, 12)       =     1.02
                                          Prob > F        =    0.4344
```

(Std. Err. adjusted for 13 clusters in n)

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
x1	-4.031638	5.348209	-0.75	0.465	-15.68439 7.621109
x2	-7.759379	9.092318	-0.85	0.410	-27.56984 12.05108
x3	-7.932472	4.96296	-1.60	0.136	-18.74583 2.88089
x4	-25.99093	70.89335	-0.37	0.720	-180.4543 128.4724
_cons	.0367635	.0236685	1.55	0.146	-.0148057 .0883327
sigma_u	.03452333				
sigma_e	.04870615				
rho	.33440243	(fraction of variance due to u_i)			

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 455.39  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 8.757  
Prob > F = 0.0119

**Table A4: Food products, beverages and tobacco (II)- population to replace X3**

```
. xtreg y1 x1 x2 dlnpop x4, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =    117
Group variable: n                     Number of groups =    13
R-sq:  within = 0.1118                Obs per group:  min =     9
      between = 0.0865                  avg           =    9.0
      overall  = 0.0088                  max           =     9
corr(u_i, Xb) = -0.6936                F(4, 12)       =     1.02
                                          Prob > F        =    0.4375
```

(Std. Err. adjusted for 13 clusters in n)

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
x1	3.893652	3.048	1.28	0.226	-2.747369 10.53467
x2	.1969173	6.414461	0.03	0.976	-13.77899 14.17283
dlnpop	-7.966934	5.006066	-1.59	0.137	-18.87421 2.940346
x4	-25.99496	70.89298	-0.37	0.720	-180.4575 128.4676
_cons	.0367886	.0237865	1.55	0.148	-.0150377 .0886149
sigma_u	.03444213				
sigma_e	.04872016				
rho	.33322728	(fraction of variance due to u_i)			

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 455.61  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 8.767  
Prob > F = 0.0119

**Table A5: Light Manufacturing (I)**

```
. xtreg y1 x1 x2 x3 x4, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0830                 Obs per group: min =    9
      between = 0.3552                 avg =           9.0
      overall = 0.0001                 max =           9

corr(u_i, Xb) = -0.7425                F(4, 12)       =    2.13
                                           Prob > F        =    0.1403

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-6.573976	6.08823	-1.08	0.301	-19.83909	6.691138
x2	3.052589	18.65288	0.16	0.873	-37.58856	43.69373
x3	-8.273735	5.802511	-1.43	0.179	-20.91632	4.368849
x4	-6.181712	6.079219	-1.02	0.329	-19.42719	7.063767
_cons	.0425391	.0276283	1.54	0.150	-.0176578	.102736
sigma_u	.03979577					
sigma_e	.06080702					
rho	.29987581	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1492.19  
Prob>chi2 = 0.0000

```
. xtserial y1 x1 x2 x3 x4
```

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 0.862  
Prob > F = 0.3715

**Table A6: Light Manufacturing (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop x4, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0825                 Obs per group: min =    9
      between = 0.3548                 avg =           9.0
      overall = 0.0001                 max =           9

corr(u_i, Xb) = -0.7416                F(4, 12)       =    2.12
                                           Prob > F        =    0.1405

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	1.69095	2.821764	0.60	0.560	-4.457145	7.839045
x2	11.34707	17.35261	0.65	0.526	-26.46103	49.15517
dlnpop	-8.302548	5.846965	-1.42	0.181	-21.04199	4.436894
x4	-6.18658	6.084441	-1.02	0.329	-19.44344	7.070279
_cons	.0425318	.0277374	1.53	0.151	-.0179027	.1029664
sigma_u	.03968709					
sigma_e	.0608237					
rho	.2986139	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1499.82  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 0.863  
Prob > F = 0.3712



**Table A9: Electricity, gas and water supply (I)**

```
. xtreg y1 x1 x2 x3 x4, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13
R-sq:  within = 0.0218                 Obs per group:  min =    9
      between = 0.7302                  avg =           9.0
      overall = 0.0052                  max =           9
corr(u_i, Xb) = -0.3756                F(4, 12)       =    1.66
                                         Prob > F        =    0.2227
                                         (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	9.873084	9.612618	1.03	0.325	-11.07101	30.81718
x2	-71.99488	63.97854	-1.13	0.282	-211.3921	67.40238
x3	-2.101343	6.856709	-0.31	0.765	-17.04083	12.83814
x4	155.3505	74.64255	2.08	0.059	-7.281657	317.9826
_cons	.0832857	.033065	2.52	0.027	.0112432	.1553282
sigma_u	.09318541					
sigma_e	.14557684					
rho	.29065074	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 20460.87  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 4.375  
 Prob > F = 0.0584

**Table A10: Electricity, gas and water supply (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop x4, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13
R-sq:  within = 0.0218                 Obs per group:  min =    9
      between = 0.7285                  avg =           9.0
      overall = 0.0050                  max =           9
corr(u_i, Xb) = -0.3725                F(4, 12)       =    1.66
                                         Prob > F        =    0.2231
                                         (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	11.96826	11.39977	1.05	0.314	-12.86971	36.80623
x2	-69.90366	58.64509	-1.19	0.256	-197.6803	57.87302
dlnpop	-2.085556	6.902023	-0.30	0.768	-17.12377	12.95266
x4	155.2001	74.76407	2.08	0.060	-7.696843	318.097
_cons	.0831761	.0331592	2.51	0.027	.0109284	.1554239
sigma_u	.09309845					
sigma_e	.14557879					
rho	.29026035	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 20495.99  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 4.379  
 Prob > F = 0.0583



**Table A13: W/sale, Retail Trade; Hotels and Rest. (I)**

```
. xtreg y1 x1 x2 x3 x4, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                    Number of groups =   13

R-sq:  within = 0.0776                Obs per group: min =    9
      between = 0.0463                avg =           9.0
      overall = 0.0134                max =           9

corr(u_i, Xb) = -0.6618                F(4, 12)       =    1.13
                                          Prob > F        =    0.3887

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-79.26534	55.95004	-1.42	0.182	-201.17	42.63933
x2	64.30553	135.1499	0.48	0.643	-230.1607	358.7718
x3	-77.174	49.84077	-1.55	0.147	-185.7677	31.41971
x4	148.0758	163.7172	0.90	0.384	-208.6334	504.785
_cons	.437367	.2184764	2.00	0.068	-.0386522	.9133862
sigma_u	.3327484					
sigma_e	.65201447					
rho	.20662995	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 83693.93  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 2.985  
Prob > F = 0.1097

**Table A14: W/sale, Retail Trade; Hotels and Rest. (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop x4, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                    Number of groups =   13

R-sq:  within = 0.0772                Obs per group: min =    9
      between = 0.0455                avg =           9.0
      overall = 0.0135                max =           9

corr(u_i, Xb) = -0.6602                F(4, 12)       =    1.12
                                          Prob > F        =    0.3924

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-2.176243	17.29201	-0.13	0.902	-39.8523	35.49982
x2	141.6776	120.5037	1.18	0.263	-120.8775	404.2327
dlnpop	-77.44784	50.24963	-1.54	0.149	-186.9324	32.03669
x4	148.1915	163.9207	0.90	0.384	-208.961	505.344
_cons	.4372917	.2194712	1.99	0.070	-.040895	.9154783
sigma_u	.33164945					
sigma_e	.65214724					
rho	.20548095	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 83905.07  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 2.987  
Prob > F = 0.1096





**Table A19: Light Manufacturing (I)**

```
. xtreg y y2 y3 y4 y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                    Number of groups =   13
R-sq:  within = 0.0856                Obs per group:  min =   10
      between = 0.1526                avg =   10.0
      overall = 0.1481                max =   10
corr(u_i, Xb) = 0.3081                F(3, 12)       =   3.20
                                       Prob > F        =  0.0620
                                       (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	67.7337	25.44937	2.66	0.021	12.28429	123.1831
y4	8.074565	7.744295	1.04	0.318	-8.798805	24.94793
y5	.0576428	.1150632	0.50	0.625	-.1930584	.3083439
_cons	2.07946	8.174465	0.25	0.804	-15.73117	19.89009
sigma_u	1.0092067					
sigma_e	.0645627					
rho	.99592405	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1744.71  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 16.848  
Prob > F = 0.0015

**Table A20: Light Manufacturing (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                    Number of groups =   13
R-sq:  within = 0.0882                Obs per group:  min =   10
      between = 0.7037                avg =   10.0
      overall = 0.6998                max =   10
corr(u_i, Xb) = 0.7189                F(4, 12)       =   3.01
                                       Prob > F        =  0.0620
                                       (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	-9.00969	8.78185	-1.03	0.325	-28.1437	10.12432
y3	56.87239	24.16504	2.35	0.036	4.221287	109.5235
lnpop	.2363639	.9073443	0.26	0.799	-1.74057	2.213297
y5	.0480182	.1346873	0.36	0.728	-.2454403	.3414767
_cons	6.299223	14.29853	0.44	0.667	-24.85459	37.45304
sigma_u	.81669045					
sigma_e	.06475375					
rho	.99375269	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1866.58  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 16.958  
Prob > F = 0.0014

**Table A21: Heavy Manufacturing (I)**

```
. xtreg y y2 y3 y4 y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.1057                Obs per group:  min =   10
      between = 0.1674                  avg   =  10.0
      overall  = 0.1669                  max   =   10
corr(u_i, Xb) = 0.0615                F(3,12)        =    2.99
                                          Prob > F        =    0.0734
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	59.57072	37.16355	1.60	0.135	-21.40171	140.5431
y4	6.699602	6.221693	1.08	0.303	-6.856302	20.25551
y5	.2342892	.1474163	1.59	0.138	-.0869033	.5554817
_cons	1.096298	6.65911	0.16	0.872	-13.41266	15.60525
sigma_u	.89807179					
sigma_e	.08422153					
rho	.99128191	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 5594.64  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 82.633  
 Prob > F = 0.0000

**Table A22: Heavy Manufacturing (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.1309                Obs per group:  min =   10
      between = 0.8486                  avg   =  10.0
      overall  = 0.8430                  max   =   10
corr(u_i, Xb) = -0.6613                F(4,12)        =    4.62
                                          Prob > F        =    0.0172
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	-9.474203	8.753762	-1.08	0.300	-28.54701	9.598606
y3	42.31306	36.79089	1.15	0.273	-37.8474	122.4735
lnpop	1.136907	1.505808	0.76	0.465	-2.143968	4.417781
y5	.1156315	.2296331	0.50	0.624	-.384696	.615959
_cons	-10.0462	23.45318	-0.43	0.676	-61.14629	41.05389
sigma_u	.5096059					
sigma_e	.08339402					
rho	.97391903	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 2493.92  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 80.735  
 Prob > F = 0.0000

**Table A23: Electricity, gas and water supply (I)**

```
. xtreg y y2 y3 y4 y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =    130
Group variable: n                     Number of groups =    13
R-sq:  within = 0.2932                 Obs per group: min =    10
      between = 0.3980                 avg =          10.0
      overall = 0.3921                 max =          10
corr(u_i, Xb) = -0.3329                F(3,12)         =    2.73
                                          Prob > F         =    0.0903
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	167.4189	134.804	1.24	0.238	-126.2938	461.1317
y4	-14.78333	34.32633	-0.43	0.674	-89.57399	60.00732
y5	1.037758	.3914173	2.65	0.021	.1849326	1.890582
_cons	13.09371	35.03086	0.37	0.715	-63.23199	89.41941
sigma_u	1.0201014					
sigma_e	.23155369					
rho	.9509988	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1077.35  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 18.549  
Prob > F = 0.0010

**Table A24: Electricity, gas and water supply (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =    130
Group variable: n                     Number of groups =    13
R-sq:  within = 0.7159                 Obs per group: min =    10
      between = 0.1065                 avg =          10.0
      overall = 0.1036                 max =          10
corr(u_i, Xb) = -0.9958                F(4,12)         =   324.03
                                          Prob > F         =    0.0000
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	-29.32887	10.99351	-2.67	0.020	-53.28167	-5.376074
y3	5.519625	104.4741	0.05	0.959	-222.1098	233.1491
lnpop	13.10346	2.226795	5.88	0.000	8.251685	17.95523
y5	.2081023	.2141403	0.97	0.350	-.2584693	.6746738
_cons	-214.0661	36.42509	-5.88	0.000	-293.4296	-134.7026
sigma_u	13.043973					
sigma_e	.14744217					
rho	.99987225	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 47214.16  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 16.988  
Prob > F = 0.0014

**Table A25: Construction (I)**

```
. xtreg y y2 y3 y4 y5, fe robust
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.5451                 Obs per group:  min =   10
      between = 0.1965                   avg =   10.0
      overall  = 0.2057                   max =   10
corr(u_i, Xb) = -0.5383                F(3, 12)       =    7.02
                                          Prob > F        =   0.0056
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	-47.18967	65.83421	-0.72	0.487	-190.6301	96.25075
y4	4.360879	10.79596	0.40	0.693	-19.1615	27.88326
y5	.9673917	.3217963	3.01	0.011	.2662577	1.668526
_cons	-2.20785	9.087008	-0.24	0.812	-22.00674	17.59104
sigma_u	1.1208527					
sigma_e	.17021871					
rho	.97745683	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 2459.00  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 1.985  
 Prob > F = 0.1842

**Table A26: Construction (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe robust
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.5669                 Obs per group:  min =   10
      between = 0.7632                   avg =   10.0
      overall  = 0.7043                   max =   10
corr(u_i, Xb) = -0.9909                F(4, 12)       =   9.63
                                          Prob > F        =   0.0010
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	3.919681	19.00681	0.21	0.840	-37.4926	45.33196
y3	-16.589	33.02559	-0.50	0.625	-88.54559	55.36759
lnpop	-3.649449	5.122851	-0.71	0.490	-14.81118	7.512284
y5	1.189026	.5922804	2.01	0.068	-.1014416	2.479495
_cons	61.0255	80.08325	0.76	0.461	-113.4609	235.5119
sigma_u	4.1685117					
sigma_e	.16682626					
rho	.99840091	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 23714.64  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 2.194  
 Prob > F = 0.1643

**Table A27: W/sale, Retail Trade; Hotels and Rest. (I)**

```
. xtreg y y2 y3 y4 y5, fe robust
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.3857                 Obs per group:  min =   10
      between = 0.0766                   avg =   10.0
      overall = 0.0897                   max =   10
corr(u_i, Xb) = -0.3298                F(3, 12)        =   22.93
                                          Prob > F         =   0.0000
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	77.53143	81.46885	0.95	0.360	-99.97394	255.0368
y4	14.88023	18.35249	0.81	0.433	-25.10642	54.86687
y5	.5660977	.1333706	4.24	0.001	.2755081	.8566873
_cons	-8.416104	18.19901	-0.46	0.652	-48.06834	31.23613
sigma_u	1.0183462					
sigma_e	.21391487					
rho	.95773915	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 15654.74  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 1.819  
 Prob > F = 0.2024

**Table A28: W/sale, Retail Trade; Hotels and Rest. (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe robust
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.4010                 Obs per group:  min =   10
      between = 0.8342                   avg =   10.0
      overall = 0.7562                   max =   10
corr(u_i, Xb) = -0.9909                F(4, 12)        =   17.13
                                          Prob > F         =   0.0001
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	-8.147728	16.05847	-0.51	0.621	-43.13613	26.84067
y3	85.05002	55.94763	1.52	0.154	-36.8494	206.9494
lnpop	-2.750414	1.638235	-1.68	0.119	-6.319822	.8189943
y5	.6594793	.1564379	4.22	0.001	.3186303	1.000328
_cons	51.43497	26.17944	1.96	0.073	-5.605128	108.4751
sigma_u	3.4565356					
sigma_e	.21217494					
rho	.99624619	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 12049.45  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 1.721  
 Prob > F = 0.2141

## Changes in employment

Y1= employment (t - (t-1))/ employment (t - 1)

X1= immigrants from EECs (t - (t-1)) / total population (t - 1)

X2= immigrants from MPCs (t - (t-1))/ total population (t - 1)

X3= native people (t - (t-1))/ total population (t - 1)

**Table A29: Mining and quarrying (I)**

```
. xtreg y1 x1 x2 x3, fe robust
```

Fixed-effects (within) regression  
Group variable: n

Number of obs = 117  
Number of groups = 13

R-sq: within = 0.0047  
between = 0.0339  
overall = 0.0002

Obs per group: min = 9  
avg = 9.0  
max = 9

corr(u\_i, Xb) = -0.3223

F(3,12) = 0.50  
Prob > F = 0.6872

(Std. Err. adjusted for 13 clusters in n)

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	2.828564	3.138674	0.90	0.385	-4.010019	9.667146
x2	-12.45903	15.24799	-0.82	0.430	-45.68155	20.76348
x3	2.131668	2.394258	0.89	0.391	-3.084972	7.348308
_cons	-.0185745	.0113436	-1.64	0.127	-.0432902	.0061412
sigma_u	.02137637					
sigma_e	.08437394					
rho	.0603161					(fraction of variance due to u_i)

Modified Wald test for groupwise heteroskedasticity  
in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 690.87  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 0.258  
Prob > F = 0.6207

**Table A30: Mining and quarrying (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop, fe robust
```

Fixed-effects (within) regression  
Group variable: n

Number of obs = 117  
Number of groups = 13

R-sq: within = 0.0048  
between = 0.0341  
overall = 0.0002

Obs per group: min = 9  
avg = 9.0  
max = 9

corr(u\_i, Xb) = -0.3234

F(3,12) = 0.51  
Prob > F = 0.6858

(Std. Err. adjusted for 13 clusters in n)

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.6957452	1.968906	0.35	0.730	-3.594132	4.985622
x2	-14.60521	15.56878	-0.94	0.367	-48.52667	19.31625
dlnpop	2.156283	2.412193	0.89	0.389	-3.099435	7.412
_cons	-.0186534	.0113866	-1.64	0.127	-.0434627	.0061559
sigma_u	.02138914					
sigma_e	.08437272					
rho	.06038544					(fraction of variance due to u_i)

Modified Wald test for groupwise heteroskedasticity  
in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 690.92  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 0.258  
Prob > F = 0.6206

**Table A31: Food products, beverages and tobacco (I)**

```
. xtreg y1 x1 x2 x3, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0063                Obs per group:  min =    9
      between = 0.0862                  avg   =   9.0
      overall  = 0.0007                  max   =    9

corr(u_i, Xb) = -0.4715                F(3,12)         =    1.54
                                           Prob > F         =    0.2553

                                           (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-1.517677	2.958406	-0.51	0.617	-7.96349	4.928136
x2	-1.261237	11.34956	-0.11	0.913	-25.98981	23.46734
x3	-1.539014	.8212806	-1.87	0.085	-3.328431	.2504024
_cons	.002266	.0037843	0.60	0.560	-.0059793	.0105113
sigma_u	.01449154					
sigma_e	.04207326					
rho	.10605427	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 230.21  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 5.408  
Prob > F = 0.0384

**Table A32: Food products, beverages and tobacco (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0063                Obs per group:  min =    9
      between = 0.0860                  avg   =   9.0
      overall  = 0.0007                  max   =    9

corr(u_i, Xb) = -0.4709                F(3,12)         =    1.53
                                           Prob > F         =    0.2562

                                           (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.0206653	2.801857	0.01	0.994	-6.084057	6.125388
x2	.2843025	11.73508	0.02	0.981	-25.28424	25.85285
dlnpop	-1.5493	.8271909	-1.87	0.086	-3.351595	.2529939
_cons	.0022878	.003798	0.60	0.558	-.0059872	.0105628
sigma_u	.01448657					
sigma_e	.04207326					
rho	.10598923	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 230.26  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 5.412  
Prob > F = 0.0383

**Table A33: Light Manufacturing (I)**

```
. xtreg y1 x1 x2 x3, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0210                 Obs per group:  min =    9
        between = 0.0083                avg =           9.0
        overall = 0.0045                max =           9

corr(u_i, Xb) = -0.4036                F(3, 12)       =    1.19
                                           Prob > F        =    0.3560

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-7.18586	7.670385	-0.94	0.367	-23.89819	9.526474
x2	18.71142	15.97737	1.17	0.264	-16.10028	53.52312
x3	-4.340812	2.360225	-1.84	0.091	-9.483301	.8016767
_cons	.0143627	.0111306	1.29	0.221	-.0098889	.0386142
sigma_u	.027731					
sigma_e	.07435501					
rho	.1221099	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 18614.72  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 0.089  
 Prob > F = 0.7708

**Table A34: Light Manufacturing (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0211                 Obs per group:  min =    9
        between = 0.0082                avg =           9.0
        overall = 0.0046                max =           9

corr(u_i, Xb) = -0.4031                F(3, 12)       =    1.18
                                           Prob > F        =    0.3572

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-2.846265	7.002979	-0.41	0.692	-18.10444	12.41192
x2	23.07242	17.26133	1.34	0.206	-14.53679	60.68163
dlnpop	-4.373204	2.383085	-1.84	0.091	-9.565499	.8190919
_cons	.01444	.0111991	1.29	0.222	-.0099607	.0388407
sigma_u	.02772328					
sigma_e	.07435409					
rho	.12205282	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 18400.14  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 0.091  
 Prob > F = 0.7687

**Table A35: Heavy Manufacturing (I)**

```
. xtreg y1 x1 x2 x3, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0315                 Obs per group:  min =    9
      between = 0.0071                 avg =           9.0
      overall = 0.0244                 max =           9

corr(u_i, Xb) = -0.1499                F(3,12)        =    0.64
                                           Prob > F        =    0.6041

                                           (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-.156907	2.681787	-0.06	0.954	-6.000019	5.686205
x2	26.8917	21.86134	1.23	0.242	-20.74006	74.52347
x3	-2.670229	2.57361	-1.04	0.320	-8.277644	2.937186
_cons	.0093407	.0116087	0.80	0.437	-.0159525	.0346339
sigma_u	.01890769					
sigma_e	.06059977					
rho	.08871353	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1789.90  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 2.037  
 Prob > F = 0.1790

**Table A36: Heavy Manufacturing (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0315                 Obs per group:  min =    9
      between = 0.0072                 avg =           9.0
      overall = 0.0245                 max =           9

corr(u_i, Xb) = -0.1487                F(3,12)        =    0.64
                                           Prob > F        =    0.6050

                                           (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	2.511371	2.857931	0.88	0.397	-3.715527	8.738268
x2	29.57119	22.92743	1.29	0.221	-20.38338	79.52576
dlnpop	-2.684163	2.595958	-1.03	0.322	-8.340269	2.971943
_cons	.0093601	.0116688	0.80	0.438	-.016064	.0347842
sigma_u	.01889535					
sigma_e	.06060056					
rho	.08860596	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1783.13  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 2.036  
 Prob > F = 0.1791

**Table A37: Electricity, gas and water supply (I)**

```
. xtreg y1 x1 x2 x3, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0115                Obs per group:  min =    9
        between = 0.1557                avg   =   9.0
        overall = 0.0267                max   =    9

corr(u_i, Xb) = -0.0320                F(3,12)        =    1.10
                                           Prob > F        =    0.3857

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	4.876119	5.035444	0.97	0.352	-6.095172	15.84741
x2	4.831504	22.11056	0.22	0.831	-43.34326	53.00627
x3	3.212526	2.087078	1.54	0.150	-1.334825	7.759878
_cons	-.0229599	.0097086	-2.36	0.036	-.0441131	-.0018067
sigma_u	.02143709					
sigma_e	.06981997					
rho	.08614855	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 11763.82  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 1.877  
Prob > F = 0.1958

**Table A38: Electricity, gas and water supply (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0115                Obs per group:  min =    9
        between = 0.1560                avg   =   9.0
        overall = 0.0267                max   =    9

corr(u_i, Xb) = -0.0307                F(3,12)        =    1.10
                                           Prob > F        =    0.3880

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	1.665826	5.007096	0.33	0.745	-9.243699	12.57535
x2	1.607537	22.26182	0.07	0.944	-46.8968	50.11187
dlnpop	3.229865	2.105609	1.53	0.151	-1.357862	7.817593
_cons	-.0229859	.0097592	-2.36	0.036	-.0442495	-.0017223
sigma_u	.02143297					
sigma_e	.06982084					
rho	.08611631	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 11767.38  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 1.877  
Prob > F = 0.1958

**Table A39: Construction (I)**

```
. xtreg y1 x1 x2 x3, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0164                 Obs per group:  min =    9
      between = 0.1262                 avg   =   9.0
      overall = 0.0045                 max   =    9

corr(u_i, Xb) = -0.3574                F(3,12)        =    3.47
                                           Prob > F        =   0.0509

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-22.35255	21.21779	-1.05	0.313	-68.58213	23.87703
x2	78.19556	42.94908	1.82	0.094	-15.38246	171.7736
x3	5.046003	6.318253	0.80	0.440	-8.720288	18.81229
_cons	.0112637	.0307797	0.37	0.721	-.0557995	.078327
sigma_u	.05216073					
sigma_e	.22322164					
rho	.05177568	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1.6e+05  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 2.328  
Prob > F = 0.1530

**Table A40: Construction (II) - population to replace X3**

```
. xtreg y1 x1 x2 dlnpop, fe robust
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0164                 Obs per group:  min =    9
      between = 0.1278                 avg   =   9.0
      overall = 0.0045                 max   =    9

corr(u_i, Xb) = -0.3561                F(3,12)        =    3.47
                                           Prob > F        =   0.0509

                               (Std. Err. adjusted for 13 clusters in n)
```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-27.38829	17.76619	-1.54	0.149	-66.09749	11.3209
x2	73.14934	46.82628	1.56	0.144	-28.87637	175.175
dlnpop	5.039588	6.373687	0.79	0.444	-8.847482	18.92666
_cons	.011381	.0309371	0.37	0.719	-.0560251	.0787872
sigma_u	.05211532					
sigma_e	.22322581					
rho	.05168839	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1.6e+05  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 2.323  
Prob > F = 0.1534

**Table A41: W/sale, Retail Trade; Hotels and Rest. (I)**

```

. xtreg y1 x1 x2 x3, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0525                Obs per group: min =    9
      between = 0.0065                avg =           9.0
      overall = 0.0123                max =           9

corr(u_i, Xb) = -0.5788                F(3,12)         =    1.46
                                          Prob > F         =    0.2759

                                          (Std. Err. adjusted for 13 clusters in n)

```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-61.15215	54.58952	-1.12	0.285	-180.0925	57.7882
x2	121.2561	141.9327	0.85	0.410	-187.9887	430.5009
x3	-69.42766	50.55852	-1.37	0.195	-179.5852	40.72989
_cons	.429306	.2427318	1.77	0.102	-.0995611	.9581731
sigma_u	.30208011					
sigma_e	.66091658					
rho	.17280585 (fraction of variance due to u_i)					

```

Modified Wald test for groupwise heteroskedasticity
in fixed effect regression model

H0: sigma(i)^2 = sigma^2 for all i

chi2 (13) = 1.1e+06
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data
H0: no first-order autocorrelation
F( 1, 12) = 33.305
Prob > F = 0.0001

```

**Table A42: W/sale, Retail Trade; Hotels and Rest. (II) - population to replace X3**

```

. xtreg y1 x1 x2 dlnpop, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   117
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0522                Obs per group: min =    9
      between = 0.0061                avg =           9.0
      overall = 0.0124                max =           9

corr(u_i, Xb) = -0.5764                F(3,12)         =    1.45
                                          Prob > F         =    0.2783

                                          (Std. Err. adjusted for 13 clusters in n)

```

y1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
x1	8.196616	17.67514	0.46	0.651	-30.3142	46.70743
x2	190.8505	129.1669	1.48	0.165	-90.58011	472.281
dlnpop	-69.6497	50.94174	-1.37	0.197	-180.6422	41.34282
_cons	.429151	.2436714	1.76	0.104	-.1017634	.9600655
sigma_u	.30100039					
sigma_e	.66103332					
rho	.17173432 (fraction of variance due to u_i)					

```

Modified Wald test for groupwise heteroskedasticity
in fixed effect regression model

H0: sigma(i)^2 = sigma^2 for all i

chi2 (13) = 1.1e+06
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data
H0: no first-order autocorrelation
F( 1, 12) = 33.323
Prob > F = 0.0001

```

## Employment

Y = employment

Y2 = immigrants from EECs / total population

Y3 = immigrants from MPCs / total population

Y4 = native population / total population

**Table A43: Mining and quarrying (I)**

```
. xtreg y y2 y3 y4, fe cluster(n)
```

Fixed-effects (within) regression  
Group variable: n

Number of obs = 130  
Number of groups = 13

R-sq: within = 0.0272  
between = 0.0000  
overall = 0.0000

Obs per group: min = 10  
avg = 10.0  
max = 10

corr(u\_i, Xb) = -0.0174

F(2,12) = 1.14  
Prob > F = 0.3518

(Std. Err. adjusted for 13 clusters in n)

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	-53.37374	50.98678	-1.05	0.316	-164.4644	57.7169
y4	-2.619496	7.659365	-0.34	0.738	-19.30782	14.06883
_cons	12.24406	7.669525	1.60	0.136	-4.466398	28.95452
sigma_u	1.1891735					
sigma_e	.09143632					
rho	.99412258					(fraction of variance due to u_i)

Modified Wald test for groupwise heteroskedasticity  
in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 4143.41  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 66.072  
Prob > F = 0.0000

**Table A44: Mining and quarrying (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop, fe cluster(n)
```

Fixed-effects (within) regression  
Group variable: n

Number of obs = 130  
Number of groups = 13

R-sq: within = 0.0808  
between = 0.8860  
overall = 0.8799

Obs per group: min = 10  
avg = 10.0  
max = 10

corr(u\_i, Xb) = -0.9877

F(3,12) = 1.80  
Prob > F = 0.2002

(Std. Err. adjusted for 13 clusters in n)

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	9.082087	7.619568	1.19	0.256	-7.519526	25.6837
y3	-27.908	62.90886	-0.44	0.665	-164.9746	109.1586
lnpop	-1.419316	1.154379	-1.23	0.242	-3.934492	1.09586
_cons	33.34977	19.28766	1.73	0.109	-8.674436	75.37398
sigma_u	2.5866155					
sigma_e	.08927053					
rho	.99881031					(fraction of variance due to u_i)

Modified Wald test for groupwise heteroskedasticity  
in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 1422.47  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 61.543  
Prob > F = 0.0000

**Table A45: Food products, beverages and tobacco (I)**

```
. xtreg y y2 y3 y4, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0459                 Obs per group:  min =   10
      between = 0.0017                 avg   =   10.0
      overall = 0.0009                 max   =   10

corr(u_i, Xb) = -0.0462                F(2,12)        =    1.51
                                           Prob > F       =    0.2610

                                           (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	35.04311	35.77766	0.98	0.347	-42.90972	112.9959
y4	-0.0622321	6.39897	-0.01	0.992	-14.00439	13.87993
_cons	11.69393	6.40355	1.83	0.093	-2.258206	25.64607
sigma_u	1.2906173					
sigma_e	.05254397					
rho	.99834525	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 751.11  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 10.166  
Prob > F = 0.0078

**Table A46: Food products, beverages and tobacco (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13

R-sq:  within = 0.1707                 Obs per group:  min =   10
      between = 0.3768                 avg   =   10.0
      overall = 0.3757                 max   =   10

corr(u_i, Xb) = -0.8967                F(3,12)        =    4.43
                                           Prob > F       =    0.0257

                                           (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	5.782799	3.163023	1.83	0.092	-1.108837	12.67444
y3	55.32822	33.09638	1.67	0.120	-16.78261	127.439
lnpop	-1.256344	.6127483	-2.05	0.063	-2.591408	.0787199
_cons	32.63267	10.23967	3.19	0.008	10.32236	54.94299
sigma_u	2.3014756					
sigma_e	.04920333					
rho	.99954315	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 175.91  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 11.580  
Prob > F = 0.0052

**Table A47: Light Manufacturing (I)**

```
. xtreg y y2 y3 y4, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0297                Obs per group: min =   10
       between = 0.0005                avg           =  10.0
       overall = 0.0001                max           =   10

corr(u_i, Xb) = -0.0344                F(2,12)         =    0.88
                                           Prob > F         =   0.4401

                                           (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	66.92723	50.54803	1.32	0.210	-43.20747	177.0619
y4	9.647693	9.198491	1.05	0.315	-10.3941	29.68948
_cons	3.681615	9.204759	0.40	0.696	-16.37383	23.73706
sigma_u	1.0436492					
sigma_e	.0957923					
rho	.99164573	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 5511.60  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 6.378  
 Prob > F = 0.0266

**Table A48: Light Manufacturing (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13

R-sq:  within = 0.2080                Obs per group: min =   10
       between = 0.9364                avg           =  10.0
       overall = 0.9271                max           =   10

corr(u_i, Xb) = -0.9974                F(3,12)         =    2.53
                                           Prob > F         =   0.1068

                                           (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	2.719082	4.381684	0.62	0.546	-6.827788	12.26595
y3	100.9967	46.40324	2.18	0.050	-10.73077	202.1007
lnpop	-2.715938	1.066738	-2.55	0.026	-5.040161	-.3917149
_cons	58.72877	17.83171	3.29	0.006	19.8768	97.58074
sigma_u	3.7695317					
sigma_e	.08691965					
rho	.99946859	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 20649.15  
 Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
 F( 1, 12) = 5.798  
 Prob > F = 0.0330

**Table A49: Heavy Manufacturing (I)**

```
. xtreg y y2 y3 y4, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13

R-sq:  within = 0.0530                Obs per group:  min =   10
        between = 0.0125                avg   =  10.0
        overall = 0.0096                max   =   10

corr(u_i, Xb) = 0.0756                F(2,12)        =   1.79
                                         Prob > F        =  0.2087

                                         (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	61.38328	33.046	1.86	0.088	-10.61776	133.3843
y4	7.803894	6.740877	1.16	0.270	-6.883215	22.491
_cons	5.121202	6.742348	0.76	0.462	-9.569112	19.81152
sigma_u	1.1271793					
sigma_e	.06530569					
rho	.9966545	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 2163.51  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 13.564  
Prob > F = 0.0031

**Table A50: Heavy Manufacturing (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13

R-sq:  within = 0.1894                Obs per group:  min =   10
        between = 0.8411                avg   =  10.0
        overall = 0.8373                max   =   10

corr(u_i, Xb) = -0.9862                F(3,12)        =   3.74
                                         Prob > F        =  0.0416

                                         (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	-0.3414215	4.350725	-0.08	0.939	-9.820836	9.137993
y3	79.9594	27.44146	2.91	0.013	20.16959	139.7492
lnpop	-1.638868	.7286853	-2.25	0.044	-3.226537	-.0511995
_cons	40.32033	12.18378	3.31	0.006	13.77415	66.86651
sigma_u	2.7336403					
sigma_e	.06068477					
rho	.99950744	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 456.75  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 13.096  
Prob > F = 0.0035

**Table A51: Electricity, gas and water supply (I)**

```
. xtreg y y2 y3 y4 y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =    130
Group variable: n                    Number of groups =    13
R-sq:  within = 0.2932                Obs per group:  min =    10
      between = 0.3980                avg           =   10.0
      overall  = 0.3921                max           =    10
corr(u_i, Xb) = -0.3329                F(3, 12)       =    2.73
                                          Prob > F        =   0.0903
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	167.4189	134.804	1.24	0.238	-126.2938	461.1317
y4	-14.78333	34.32633	-0.43	0.674	-89.57399	60.00732
y5	1.037758	.3914173	2.65	0.021	.1849326	1.890582
_cons	13.09371	35.03086	0.37	0.715	-63.23199	89.41941
sigma_u	1.0201014					
sigma_e	.23155369					
rho	.9509988	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 41722.91  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 25.396  
Prob > F = 0.0003

**Table A52: Electricity, gas and water supply (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe cluster(n)
Fixed-effects (within) regression      Number of obs   =    130
Group variable: n                    Number of groups =    13
R-sq:  within = 0.7159                Obs per group:  min =    10
      between = 0.1065                avg           =   10.0
      overall  = 0.1036                max           =    10
corr(u_i, Xb) = -0.9958                F(4, 12)       =   324.03
                                          Prob > F        =   0.0000
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	-29.32887	10.99351	-2.67	0.020	-53.28167	-5.376074
y3	5.519625	104.4741	0.05	0.959	-222.1098	233.1491
lnpop	13.10346	2.226795	5.88	0.000	8.251685	17.95523
y5	.2081023	.2141403	0.97	0.350	-.2584693	.6746738
_cons	-214.0661	36.42509	-5.88	0.000	-293.4296	-134.7026
sigma_u	13.043973					
sigma_e	.14744217					
rho	.99987225	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 2832.38  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data  
H0: no first-order autocorrelation

F( 1, 12) = 25.580  
Prob > F = 0.0003

**Table A53: Construction (I)**

```
. xtreg y y2 y3 y4 y5, fe robust
Fixed-effects (within) regression      Number of obs   =    130
Group variable: n                     Number of groups =    13
R-sq:  within = 0.5451                 Obs per group:  min =    10
      between = 0.1965                   avg =    10.0
      overall  = 0.2057                   max =    10
corr(u_i, Xb) = -0.5383                F(3,12)         =    7.02
                                          Prob > F         =    0.0056
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	-47.18967	65.83421	-0.72	0.487	-190.6301	96.25075
y4	4.360879	10.79596	0.40	0.693	-19.1615	27.88326
y5	.9673917	.3217963	3.01	0.011	.2662577	1.668526
_cons	-2.20785	9.087008	-0.24	0.812	-22.00674	17.59104
sigma_u	1.1208527					
sigma_e	.17021871					
rho	.97745683	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 4288.57  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 1.062  
Prob > F = 0.3231

**Table A54: Construction (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe robust
Fixed-effects (within) regression      Number of obs   =    130
Group variable: n                     Number of groups =    13
R-sq:  within = 0.5669                 Obs per group:  min =    10
      between = 0.7632                   avg =    10.0
      overall  = 0.7043                   max =    10
corr(u_i, Xb) = -0.9909                F(4,12)         =    9.63
                                          Prob > F         =    0.0010
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	3.919681	19.00681	0.21	0.840	-37.4926	45.33196
y3	-16.589	33.02559	-0.50	0.625	-88.54559	55.36759
lnpop	-3.649449	5.122851	-0.71	0.490	-14.81118	7.512284
y5	1.189026	.5922804	2.01	0.068	-.1014416	2.479495
_cons	61.0255	80.08325	0.76	0.461	-113.4609	235.5119
sigma_u	4.1685117					
sigma_e	.16682626					
rho	.99840091	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 8134.55  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 0.848  
Prob > F = 0.3752

**Table A55: W/sale, Retail Trade; Hotels and Rest. (I)**

```
. xtreg y y2 y3 y4 y5, fe robust
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.3857                 Obs per group: min =   10
      between = 0.0766                  avg           =  10.0
      overall  = 0.0897                 max           =   10
corr(u_i, Xb) = -0.3298                F(3,12)         =   22.93
                                          Prob > F         =   0.0000
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	(dropped)					
y3	77.53143	81.46885	0.95	0.360	-99.97394	255.0368
y4	14.88023	18.35249	0.81	0.433	-25.10642	54.86687
y5	.5660977	.1333706	4.24	0.001	.2755081	.8566873
_cons	-8.416104	18.19901	-0.46	0.652	-48.06834	31.23613
sigma_u	1.0183462					
sigma_e	.21391487					
rho	.95773915	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 65624.57  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 1046.216  
Prob > F = 0.0000

**Table A56: W/sale, Retail Trade; Hotels and Rest. (II) - population to replace Y4**

```
. xtreg y y2 y3 lnpop y5, fe robust
Fixed-effects (within) regression      Number of obs   =   130
Group variable: n                     Number of groups =   13
R-sq:  within = 0.4010                 Obs per group: min =   10
      between = 0.8342                  avg           =  10.0
      overall  = 0.7562                 max           =   10
corr(u_i, Xb) = -0.9909                F(4,12)         =   17.13
                                          Prob > F         =   0.0001
                                          (Std. Err. adjusted for 13 clusters in n)
```

y	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
y2	-8.147728	16.05847	-0.51	0.621	-43.13613	26.84067
y3	85.05002	55.94763	1.52	0.154	-36.8494	206.9494
lnpop	-2.750414	1.638235	-1.68	0.119	-6.319822	.8188943
y5	.6594793	.1564379	4.22	0.001	.3186303	1.000328
_cons	51.43497	26.17944	1.96	0.073	-5.605128	108.4751
sigma_u	3.4565356					
sigma_e	.21217494					
rho	.99624619	(fraction of variance due to u_i)				

Modified Wald test for groupwise heteroskedasticity in fixed effect regression model

H0:  $\sigma(i)^2 = \sigma^2$  for all i

chi2 (13) = 2.7e+06  
Prob>chi2 = 0.0000

Wooldridge test for autocorrelation in panel data

H0: no first-order autocorrelation  
F( 1, 12) = 955.471  
Prob > F = 0.0000